

# Workshop BCB: Macro com agentes heterogêneos

## Aula 9 e 10: Calibration and Estimation in HA Models

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# Introduction

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- We have to choose the parameters to our model.
- How to choose parameters?
- What moments are important?
- What is the difference between RA and HA models?
- **Rules of the Game:** Some parameters are fixed outside the model, others are calibrated to match some micro moments, and others are estimated using some macro moments.

# Rules of the Game

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- Calibrate parameters before solving the model:
  - ▶ **Key:** Earnings process, progressive taxation/transfers.
- Calibrate in the steady state to micro moments:
  - ▶ **Key:** Wealth distribution, marginal propensities to consume.
- Estimate using macro moments (time-series).
  - ▶ **Key:** Inequality, variance of earnings growth (risk) over the cycle.
- Conceptually not very different relative to representative agent model, but the set of “micro moments” are much larger.
- I will focus on the important parameters for the HA part, the “supply” side of the model (NKPC, production function, etc) is the same across models.

# Income Process

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- The first, and perhaps one of the most important parameters in HA models are the parameters of the earnings process.
- In the simplest version, the stochastic process is given by:

$$y_{it} = \rho y_{it-1} + \varepsilon_{it}, \quad \varepsilon_t \sim N(0, \sigma^2) \quad (1)$$

- To estimate this process you need panel data of income of at least 2-periods so you have enough information on the persistence ( $\rho$ ) and inequality ( $\sigma^2$ ).
- As it will become clear later, this is a simple process, more involved processes will require more information - either higher moments or longer time series.

# Intuition Income Process

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- Note that there are two settings of moments informative about the process, the variance of earnings in levels,  $V(y_{it})$ ,  $V(y_{it-1})$  and in growth  $V(\Delta y_{it})$ , where  $\Delta y_{it} = y_{it} - y_{it-1}$
- Taking the variance in equation (1)

$$V(y_{it}) = \rho^2 V(y_{it-1}) + \sigma^2$$

- Subtracting  $y_{t-1}$  in both sides and taking the variance in equation (1):

$$y_{it} - y_{t-1} = (\rho - 1)y_{it-1} + \varepsilon_{it},$$

$$V(\Delta y_{it}) = (\rho - 1)^2 V(y_{it-1}) + \sigma^2$$

# Basic Income Process

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- You have 2 equations, 2 unknowns  $(\rho, \sigma^2)$ , and three moments  $V(y_{it})$ ,  $V(y_{it-1})$ ,  $V(\Delta y_{it})$   
 $\Rightarrow$  overidentified system:

$$\begin{aligned}V(y_{it}) &= \rho^2 V(y_{it-1}) + \sigma^2 \\V(\Delta y_{it}) &= (\rho - 1)^2 V(y_{it-1}) + \sigma^2\end{aligned}$$

- If you assume the system is stationary you can use either  $V(y_{it-1})$  or  $V(y_{it})$  (in infinite horizon).
- In life-cycle models, you could use the extra moment to identify initial heterogeneity.
- Note that you can substitute  $V(\Delta y_{it})$  by the autocovariance  $C(y_{it}, y_{it-1})$ .

# Transitory-persistent Process

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- A popular alternative is to model the earnings process as the sum of transitory and a persistent component:

$$y_{it} = z_{it} + \varepsilon_{it}$$

$$z_{it} = \rho z_{it-1} + \eta_{it}$$

where  $\varepsilon_{it} \sim N(0, \sigma_\varepsilon^2)$  is the shock of the transitory component, and  $\eta_{it} \sim N(0, \sigma_\eta^2)$  the shock of the persistent component.

- ▶ **Transitory:** Bonus, health shocks, short unemployment spells
  - ▶ **Persistent:** Promotions, unemployment spells with scarring effects.
- Persistent shocks matter more for welfare and savings behavior.

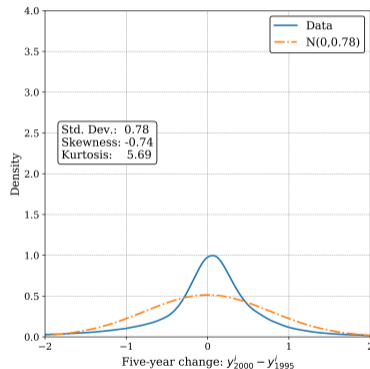
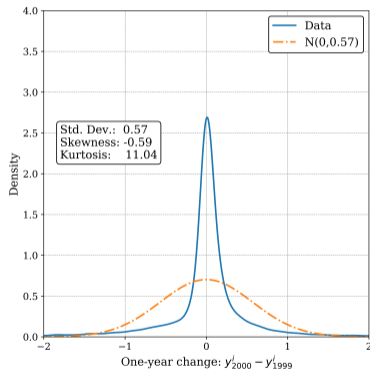
# Transitory-persistent Process

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- The transitory-persistent provides better fit and captures the income dynamics of longer horizon.
- Requires **at least four periods** of panel data.
- It can still be discretized using the usual methods, but the state space increases fast.
- Identification requires the autocovariance matrix of earnings (in growth rate or in levels). Estimation usually done using minimum distance/GMM.
- More information on the econometric identification: Guvenen (RED, 2009) identification in levels; Blundell, Pistaferri and Preston (AER, 2008) identification in growth.

# Higher moments of earnings growth

- Guvenen et al (ECTA, 2021) emphasizes the role of higher-moments, non-linearities and age-dependence of earnings growth.



- This is also true in Brazil.

# Stochastic process with higher moments

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- Higher moments give additional incentives for precautionary savings. We can specify the earnings processes with higher-moments:

$$y_{it} = z_{it} + \varepsilon_{it},$$

$$z_{it} = z_{it-1} + \eta_{it},$$

$$\eta_{it} \sim \begin{cases} N(\mu_{\eta,1}, \sigma_{\eta,1}^2) & \text{with prob. } p_{\eta} \\ N(\mu_{\eta,2}, \sigma_{\eta,2}^2) & \text{with prob. } 1 - p_{\eta} \end{cases}$$

$$\varepsilon_{it} \sim \begin{cases} N(\mu_{\varepsilon,1}, \sigma_{\varepsilon,1}^2) & \text{with prob. } p_{\varepsilon}, \\ N(\mu_{\varepsilon,2}, \sigma_{\varepsilon,2}^2) & \text{with prob. } 1 - p_{\varepsilon}. \end{cases}$$

where the shocks are drawn from a mixture of normals. Other distributions are also possible.

## Stochastic process with higher moments

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- Still requires long panel data and specially you must feed **higher moments** of the distribution in the estimation of the extra parameters  $(p_\eta, p_\varepsilon, \dots)$ .
- Luckily, the moments of the earnings growth distribution (for Brazil) are available in the GRID project: <https://www.grid-database.org/>.
- You must be careful and think whether your moments identify the higher moments.
- Estimation is usually done through simulated methods of moments (SMM). It is slow, but it is done outside of the model.
- Discretization is not trivial, but can be done relatively fast using simulation methods. The reference is: DiNardi et al (JEEA, 2020).

## Other approaches and Extra Issues

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- You can combine other shocks in the stochastic process to capture different dimensions not captured by income:
  - ▶ **Unemployment:** with some exogenous probability the agent becomes unemployed.
  - ▶ **Superstar shock/entrepreneurs:** with some exogenous probability the agent becomes an entrepreneur (Castañeda et al, 2003; Bayer and Luetticke many papers).
- What other features could be incorporated?
- **Business cycles:** There is a large literature on the cyclicalities of risk, including higher moments.
  - ▶ HA literature knows that this matters for precautionary savings and consumption (McKay, JME, 2017) but still relatively unexplored in HANK (exception is Bayer et al. ECTA, 2019).

# Progressive Taxation

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- Idiosyncratic shocks imply earnings inequality. Progressive taxation matters, since it redistribute from the top to the bottom: changes wealth distribution, MPCs, etc.
- Suppose the tax function has the following form:

$$y_i^n = F(y_i),$$

where  $y_i^n$  is net income and  $y_i$  is gross income.

- What function should we use? Two approaches:
  - ▶ Log-linear form;
  - ▶ Brackets;

# Log-linear Form

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- A functional form that captures progressivity (See Benabou (2002), Heathcote et al. (2017)):

$$T(y) = y - \tau_1 y^{1-\tau_2} \quad \text{where } y \text{ is the individual gross labor income.}$$

- ▶  $\tau_2$  gives the degree of progressivity, i.e. it measures the elasticity of posttax to pretax income.
  - ▶ Given  $\tau_2$ ,  $\tau_1$  shifts the tax function and determines the average level of taxation in the economy.
- This implies that map from gross income to net income is:

$$y_i^n = F(y_i) = y_i - T(y_i) = \tau_1 y_i^{1-\tau_2}$$

- Parameters can be easily estimated in regressing  $\log y_i^n$  on  $\log y_i$ .

# Log-linear Form

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- The tax is progressive if the ratio of marginal to average tax rates is larger than 1 for every level of income.
  - ▶  $\tau_2 = 1$ : full redistribution  $\Rightarrow T(y) = y - \tau_1$ .
  - ▶  $0 < \tau_2 < 1$ : progressivity  $\Rightarrow T'(y) > \frac{T(y)}{y}$ .
  - ▶  $\tau_2 = 0$ : no redistribution  $\Rightarrow T'(y) = \frac{T(y)}{y} = 1 - \tau_1$ .
  - ▶  $\tau_2 < 0$ : regressivity  $\Rightarrow T'(y) < \frac{T(y)}{y}$ .
- Break-even income:  $y_{be} = \tau_1^{\frac{1}{\tau_2}}$ .
  - ▶ If  $y_i > y_{be}$ ,  $i$  is a taxpayer.
  - ▶ If  $y_i < y_{be}$ ,  $i$  receives a transfer.

# Progressive Taxation

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- Log-linear:
  - ▶ **Good:** Flexible; Easy to estimate if you have the data.
  - ▶ **Bad:** Cannot account for specific marginal rates; Cannot be estimated if you do not have gross and net income for the same  $i$  (in the US they input using TAXSIM).
- Alternative: replicate the actual tax system in the function  $F$ .
- Include brackets of all marginal rates, but also possible transfers. **Brackets:**
  - ▶ **Good:** Account for top marginal rates. Very flexible.
  - ▶ **Bad:** How to model the **entire** transfer system? What to include and what to leave out?

# Wealth Distribution

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- Getting a “correct” wealth distribution was at the core of the early literature of heterogeneous agents.
- Early papers → getting the top right
- Various approaches (see DiNardi and Fella, RED, 2017):
  - ▶ Correct income process;
  - ▶ Preference heterogeneity;
  - ▶ Life-cycle motives: bequest, human capital, health shocks;
  - ▶ Entrepreneurship.
  - ▶ Heterogeneity (and shocks) in  $r_t$ .
- HANK papers → getting the bottom right → getting the right MPC (core mechanism of transmission of aggregate shocks).

# Wealth Distribution

- Which moments to target?
- **Example:** Kaplan, Moll and Violante:

TABLE 5

	Data	Model	Moment	Liquid wealth		Illiquid wealth	
				Data	Model	Data	Model
Mean illiquid assets	2.92	2.92	Top 0.1 percent share	17	2.3	12	7
Mean liquid assets	0.26	0.23	Top 1 percent share	47	18	33	40
Frac. with $b = 0$ and $a = 0$	0.10	0.10	Top 10 percent share	86	75	70	88
Frac. with $b = 0$ and $a > 0$	0.20	0.19	Bottom 50 percent share	-4	-3	3	0.1
Frac. with $b < 0$	0.15	0.15	Bottom 25 percent share	-5	-3	0	0
Gini coefficient				0.98	0.86	0.81	0.82

*Notes:* Left panel: moments targeted in calibration and reproduced by the model. Means are expressed as ratios to annual output. Right panel: statistics for the top and bottom of the wealth distribution not targeted in the calibration.

*Source:* SCF 2004

# Calibrating the Wealth Distribution

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- Early approach: permanent heterogeneity in  $\beta$  (Krussel-Smith, 1998).
- For instance, suppose:  $\beta \in [\bar{\beta} - \epsilon, \bar{\beta} + \epsilon]$ .
- Discretize the space of  $\beta$  with uniform probability (Krueger, Mitman and Perri, 2016).
- You can also calibrate the beta of each group  $g$  individually:  $\beta$  targeting specific moments of the percentiles of the wealth distribution.
- Then  $\bar{\beta}$  to match wealth-to-income ratio / avg. interest rate / avg. level of liquid asset.

# Calibrating the Wealth Distribution

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- Use the portfolio adjustment cost function (Kaplan and Violante ECTA 2014, Kaplan, Moll, Violante).
- Recall in KMV:  $\chi(d, a) = \chi_0|d| + \chi_1|d/a|^{\chi_2}$
- Choose  $(\rho, \kappa, \chi_1, \chi_2, \chi_3)$  to match fraction of individuals at the borrowing constraint, with negative wealth and mean liquid/illiquid assets.
- **Bayer, Born and Luetticke**: use wedge of interest rate between deposits and debt, and probability of portfolio rebalance to match ratio of liquid-illiquid, share of borrowers.

- **Even better**  $\Rightarrow$  we can also target the aggregate MPC.
- Auclert, Rognlie, Straub (2023): target the MPC over the wealth distribution.
- **Problem:** In Brazil there are little data on wealth, MPC is even worse.
- What data there is in the BCB to calibrate these models?
  - ▶ Share borrowers?
  - ▶ Avg. value of liquid assets?
  - ▶ Fluctuation in credit card?

- Comparing SW with Bayer and Luetticke...
- ... Acharya, Chen, del Negro, Dogra, Goyal, Matlin, Lee, Sarfati, Sengupta (Estimating HANK for Central Banks) write:

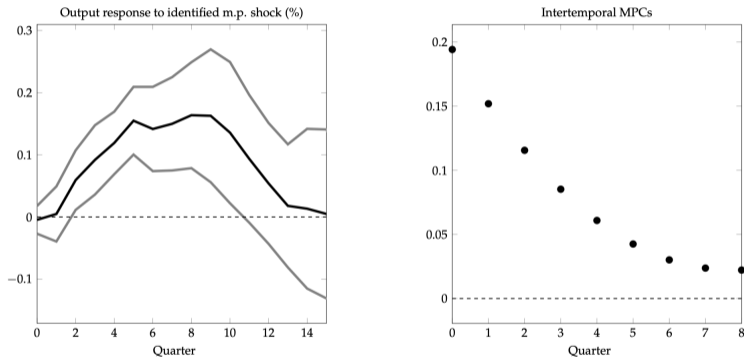
“We find that HANK’s accuracy for real activity variables is **notably inferior to that of SW**. The results for consumption are **disappointing**...”
- Why HANKs are still far from medium and large-scale DSGE models?

- Incipient literature estimating HANK models.
- Few papers:
  - ▶ Auclert, Rognlie and Straub (R&R AER, 2023): Estimation using Bayesian methods and matching IRF.
  - ▶ Bayer, Born, Luetticke (cond. accepted AER, 2023): Estimation using Bayesian methods.
  - ▶ Hagedorn, Manovskii and Mitman (WP, 2019): matching IRF.
- Usually they try to keep the supply side as close as possible to Smets and Wouters (2007) and Cristiano, Eichenbaum and Evans (2005).

- Where  $HA \neq RA \Rightarrow$  consumption function.
- Auclert, Rognlie and Straub (R&R AER, 2023):
  - ▶ Standard HA models can match the **micro jumps** in consumption out of transitory income changes (i.e., MPCs)...
  - ▶ ... but cannot match the **macro humps** observed in the aggregate consumption IRF.
- **Crucial:** the trick used in RA models to get the macro hump, **habit formation**, cannot be used in HA models.

# Macro Humps vs Micro Jumps

Figure 1: Macro Humps, Micro Jumps.



*Note.* Left panel shows the impulse response of output to a [Romer and Romer \(2004\)](#) shock, estimated with a [Jordà \(2005\)](#) projection; see section 4.2 for details. Right panel shows the consumption response to a one-time unanticipated increase in average labor incomes; estimated by [Fagereng, Holm and Natvik \(2018\)](#) using Norwegian administrative data; interpolated to quarterly data using cubic interpolation on the cumulative spending response.

# Why Not Habit Formation?

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- Why we cannot use habit formation in HA models?
- **Smets and Wouters:**  $u(c - \gamma C_-)$ , where  $C_-$  is avg consumption of previous period.
  - ▶ Usually:  $\gamma = 0.6$ .
  - ▶ HA model would need many agents below  $0.6C_-$ , implying infinite MUC.
- **Cristiano, Eichenbaum and Evans:**  $u(c - \gamma c_-)$ , where  $c_-$  is the agent's own consumption of previous period.
  - ▶ Substantially lower MPCs.
  - ▶ Implies increasing iMPCs, the opposite of the data.

# Sticky Information

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- Proposed solution: **Sticky Information** (Mankiw and Reis, 2002, 2007).
- Individuals update their expectations about the aggregate state of the economy with prob.  $1 - \theta$ .
  - ▶ Assume  $r_t$  and  $Y_t$  are the aggregate variables that follow a stochastic process.
  - ▶ Does not affected expectations of idiosyncratic shock.
- Recursive problem (not showing the budget constraint):

$$V_t(b, s, k) = \max_{c, b'} u(c) + \beta \mathbb{E}_{t-k} [\theta V_{t+1}(b', s', k+1) + (1 - \theta) V_{t+1}(b', s', 0)]$$

where  $b$  is liquid assets,  $s$  earnings process and  $k$  **the last period the agent updated its information set.**

# Sticky Information

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- Idiosyncratic shocks are functions of the aggregate outcome:  $s_t Y_t$ , so agents always observe  $Y_t$  and  $r_t$  and borrowing constraint is not affected.
- The only channel is through the expectations:  $\mathbb{E}_{t-k}$ .
- Then, we get what we want:
  - ▶ Intertemporal MPCs are unchanged since unanticipated income shock does not change future income.
  - ▶ Slow adjustment of expectations allows us to model hump-shaped impulse responses.

# Estimated “Inattentive” HANK Model

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## Full Model:

- HA with two-assets, **sticky information on the value of his illiquid account**.
- **Permanent heterogeneity**: six groups ex-ante heterogeneous in  $\beta$ , avg. income, adj. cost of illiquid asset (matching avg. illiquid asset in a group).
- Agents save in illiquid asset when they expect to be high, and dissave when they expect to be low.
  - ▶ Get the delayed aggregate consumption response (the hump).
- Rest of the model is standard:
  - ▶ Sticky-prices and wages, investment adj. cost, inertial Taylor rule, non-arbitrage asset pricing, fiscal rules with debt adjustment.

# Calibration: Heterogeneity in Illiquid Assets

Table 1: Calibrating permanent household heterogeneity

Household group $g$	1	2	3	4	5	6
Population share ( $\mu_g$ )	Bottom 50%	Next 20%	Next 10%	Next 10%	Next 5%	Top 5%
Illiquid asset share	2.7%	7.0%	7.0%	13.0%	12.2%	58.0%
Labor income share	26.7%	18.3%	10.8%	14.4%	11.0%	18.8%
Discount factors (p.a.)	0.905	0.919	0.933	0.946	0.950	0.975

- Little evidence that MPC varies in illiquid asset distribution.
- Calibrate so **each group** gets the same aggregate MPC:  $\sum_{s=0}^3 \left( \frac{1}{1+r} \right)^s \frac{\partial C_s}{\partial Y_0} = 0.55$ .
- This gives the “micro jump”.

# Matching Impulse Response Functions

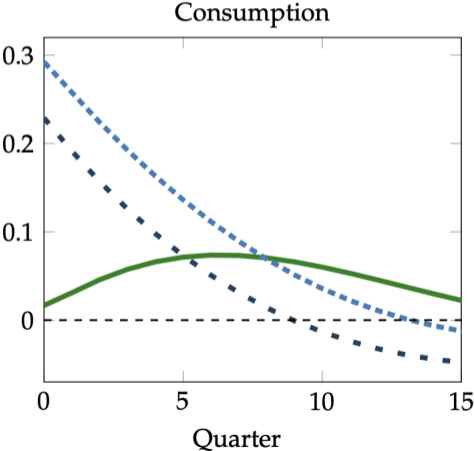
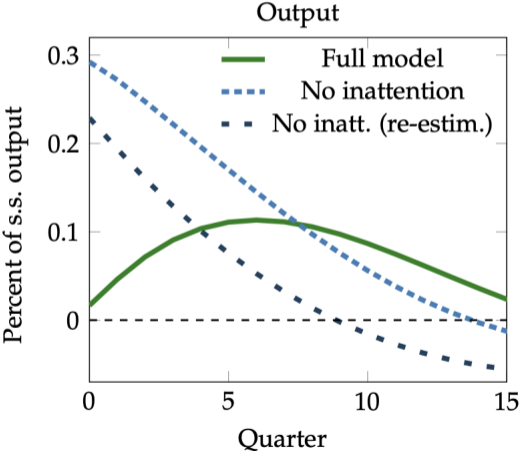
- Follow CEE (2005) and match the IRF of  $Y_t, I_t, N_t, P_t, W_t, i_t$  to identified monetary policy shocks.
- Minimize the distance between the model and the data to estimate:

**Panel B: Estimated parameters**

Parameter		Value	std. dev.
$\theta$	Household inattention	0.935	(0.01)
$\phi$	Investment adj. cost parameter	9.639	(2.428)
$\zeta_p$	Calvo price stickiness	0.926	(0.012)
$\zeta_w$	Calvo wage stickiness	0.899	(0.016)
$\rho^m$	Taylor rule inertia	0.890	(0.01)
$\sigma^m$	Std. dev. of monetary shock	0.057	(0.005)

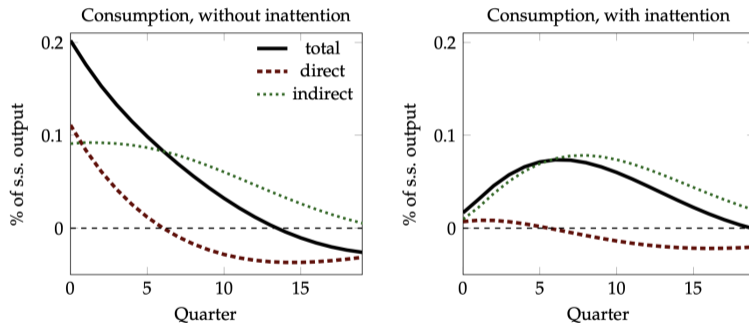
# Model Comparison 1

Figure 4: Impulse responses with and without inattention



# Direct and Indirect

Figure 6: Decomposition of consumption



- Inattention dampens the indirect effect!

# Estimation of HANK Models

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- **Auclert et al:** estimate the inattentive HANK model and highlight the role for investment as a transmission mechanism of monetary policy and of the sources of business cycles.
- **Bayer, Born, Luetticke (2023):** Exploit one advantage of the HANK model: add new data and new shocks.
- **New Data:**
  - ▶ Yearly cross-sectional information on wealth and income shares at the top 10%;
  - ▶ Time series of progressivity;
  - ▶ Income risk estimates;
- **Shocks:**
  - ▶ Income risk;
  - ▶ Progressivity of the income tax.
- Their model broadly reproduces observed US inequality dynamics.

# Estimation using the SSJ

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- **Boppart et al**: You can use the IRF and simulate the model!
- **SSJ**: You can use the IRF and compute the variance-covariance matrix using a analytical formula!
  - ▶ The  $MA(\infty)$  representation of the exogenous shock is given by:

$$dZ_t = \sum_{s=0}^{\infty} M_s^Z \epsilon_{t-s}$$

- ▶ Any endogenous variable is also  $MA(\infty)$ :

$$dX_t = \sum_{s=0}^{\infty} M_s^{X|Z} \epsilon_{t-s}$$

where the GE Jacobian,  $G$ :  $M^{X|Z} = G^{X,Z} M^Z$

- ▶ You can compute the **Variance-Covariance Matrix**:

$$\text{Cov}(dX_t, dY_{t'}) = \sum_{s=0}^{\infty} M_s^{X|Z} (M_{s+t'-t}^{Y|Z})'$$

# Likelihood Function

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- Let  $\mathbf{Y}$  be data. Assuming Gaussian innovation, we can use the V-COV-V matrix,  $\mathbf{V}$ , given a set of parameters  $\theta$  to calculate the log-likelihood of  $\theta$ :

$$\mathcal{L}(\mathbf{Y}, \theta) = -\frac{1}{2} \log \det \mathbf{V}(\theta) - \frac{1}{2} \log \det \mathbf{Y}'\mathbf{V}(\theta)^{-1}\mathbf{Y}$$

- No need for the Kalman filter.
- The costly step is to perform a Cholesky decomposition of  $\mathbf{V}$ , to get  $\mathbf{Y}'\mathbf{V}(\theta)^{-1}\mathbf{Y}$  and  $\log \det \mathbf{V}$ .
- In practice, the likelihood function has to be evaluated many times.
  - ▶ Re-use some Jacobians (or all!) if the estimated parameters do not affect the steady state.
  - ▶ Costly part is to compute the Jacobian of HA block.